Design Drives Discovery in Causal Learning

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Abstract
We assessed whether an artifact’s design can facilitate recognition of abstract causal rules. In Experiment 1, 152 three-year-olds were presented with evidence consistent with a relational rule (i.e., pairs of same or different blocks activated a machine) using two differently designed machines. In the standard-design condition, blocks were placed on top of the machine; in the relational-design condition, blocks were placed into openings on either side. In Experiment 2, we assessed whether this design cue could facilitate adults’ (N = 102) inference of a distinct conjunctive cause (i.e., that two blocks together activate the machine). Results of both experiments demonstrated that causal inference is sensitive to an artifact’s design: Participants in the relational-design conditions were more likely to infer rules that were a priori unlikely. Our findings suggest that reasoning failures may result from difficulty generating the relevant rules as cognitive hypotheses but that artifact design aids causal inference. These findings have clear implications for creating intuitive learning environments.

Keywords
causality, cognitive development, reasoning, inference, open data

Received 6/24/19; Revision accepted 10/14/19

How do we infer the causal rules that govern our everyday experiences? When reasoning about causal relationships among objects and events, learners must engage in a mental search to select the most likely explanation for their observations. For example, prior to activating a novel appliance, you might consider several hypothetical interventions: flipping the on/off switch, pressing the reset button on the circuit breaker, or perhaps using the switch and the button together to activate the device. We seem to effortlessly reason about the world, guided by our prior beliefs and experiences. But causal inference—drawing conclusions about causes by observing the occurrence of effects—is almost always undetermined; there could be infinite hypotheses to consider, and the data from our experience is not sufficient to constrain this space.

How might learners evaluate such a vast array of possible causes? Recent work suggests that learners may operate rationally, despite generating only a subset of the most likely causes to evaluate (Bonawitz & Griffiths, 2010). The specific set considered in the context of a particular problem may depend on a variety of factors, including their probability, their relevance, priming, and so forth (e.g., Dougherty & Hunter, 2003; Klein, 1993; Koehler, 1994; Schunn & Klahr, 1995). Even young children are sensitive to input that constrains the hypotheses they consider, including information about the problem, how the data were sampled, who generated the evidence, and why (e.g., Bonawitz et al., 2011; Butler & Markman, 2012; Gergely, Bekkering, & Kiraly, 2002; Walker, Lombrozo, Legare, & Gopnik, 2014). Accordingly, any input that changes a learner’s prior expectations about the most likely causal relationships among events can influence the hypotheses privileged. Here, we propose an
environmental cue that has not yet been examined in this context: an object's visible design. If learners use information about design to constrain the hypotheses they consider, the physical features of an artifact may change the salience of various causes, influencing learning and discovery.

Although design has not been specifically examined in the case of causal learning, there are reasons to expect that the physical environment influences causal inference. Indeed, all artifacts include some element of design, and we use these cues to infer their function. For example, if a door has no handle, we infer that we should push, because otherwise a handle would have been added. Norman (1988) included such constraints as one of several principles of good design that impact reasoning about the intended use of objects. An extensive literature has explored the ways in which subtle environmental influences, or “nudges,” have disproportional effects on choice (Thaler & Sunstein, 2008), impacting hygiene (Holland, Hendriks, & Aarts, 2005), energy use (Allcott & Mullainathan, 2010), and health (Thorndike, Sonnenberg, Riis, Barraclough, & Levy, 2012).

Other applied research has examined whether design can change the way we learn in select contexts. Museum designers have used exhibit access, visibility, and affordances to encourage exploration, engagement, and understanding (e.g., adding a knob suggests an object can be moved; Allen, 2004; Shin, Park, & Kim, 2014). We went beyond this applied work to consider whether similar cues can influence the salience of certain causal rules.

To illustrate how the design of an object might impact beliefs about its causal structure, consider a novel appliance that has two cords. Although you may have a strong prior belief that appliances require only a single power source, you might instead form a hypothesis that both cords must be connected; otherwise, why would the design include a second cord? This example demonstrates an even more general assumption—that an object’s features are relevant to its function—constraining the hypotheses that are generated about its mechanism (Norman, 1988).

Support for this proposal can be found in the literature examining reasoning about artifacts. Specifically, both children and adults adopt a “design stance,” viewing features of artifacts as reflective of that object’s kind, function, and intended use (Kelemen, 1999; Kelemen & Carey, 2007; Malt & Johnson, 1992). In addition, both preschoolers and adults privilege efficient design (selecting an object with a single feature for a single purpose over one with superfluous features; Kelemen, Seston, & Saint Georges, 2012) and map the quantity and diversity of object functions to infer the complexity of its internal mechanism (Ahl & Keil, 2016). Preschoolers also map the type of effect produced (i.e., discrete vs. continuous) to the mechanism that produced it (a binary on/off switch vs. a dial), providing evidence that even children relate the physical structure of an object’s mechanism to its effect (Magid, Sheskin, & Schulz, 2015).

In this prior work, learners made inferences about the design of objects, given information about functions. Here, we asked whether they could perform a more challenging task—infer an unlikely causal rule given an object’s design. In two experiments, we tested the novel prediction that manipulating the physical structure of an object would lead learners to privilege certain types of causes. We first presented a case in which 3-year-olds (Experiment 1) typically fail to privilege certain types of causes. We then presented a different case in which adults (Experiment 2) typically fail to infer an unusual conjunctive-causation rule, in which the combination of two causes produces an effect. In both paradigms, we presented each learner with one of two machines and assessed whether design differences facilitated identification of the relevant rule.

**Experiment 1**

In Experiment 1, we presented 3-year-olds with a relational-reasoning task that they systematically fail at this age (Walker, Bridgers, & Gopnik, 2016). In this task, children are introduced to a novel toy that plays music when certain pairs of blocks are placed on top. In past research, when 3-year-olds were provided with evidence that the toy is activated by the abstract relation between blocks (i.e., whether pairs are the same or different), rather than by object kinds (i.e., blocks of a particular shape or color), they failed to make the correct inference. Notably, younger children (18- to 30-month-olds) successfully infer same–different relations in the identical task, suggesting that the later decline is due to a failure to spontaneously generate relational hypotheses, not a lack of competence (Carstensen et al., 2019; Walker & Gopnik, 2014, 2017). This tendency likely results from a learned bias that temporarily privileges the role of objects over the relations between them, leading to a U-shaped developmental trajectory (in the United States1). This domain therefore provides a case study to explore whether object design influences hypothesis generation.

To assess this, we made one small modification to the original task: Rather than placing pairs of blocks on top of the machine, the experimenter inserted the blocks into transparent openings on either side. If the learner treats object design as relevant to causal inference, these features might suggest an affordance: that the machine activates because of the combination of blocks. This may raise the possibility that the relation between blocks is relevant.

On the other hand, 3-year-olds’ well-documented failure to spontaneously privilege abstract relations
suggests a strong prior for object-based hypotheses at this age. To correctly infer the relation in this case, children must integrate information about the object’s design with their prior beliefs about likely causes, taking into account why design is relevant and weighing this more heavily than their prior commitments. That said, if children are indeed sensitive to the design of the learning context, this cue may inform the types of hypotheses that are privileged, reducing their tendency to prioritize object-based causes.

**Method**

**Participants.** A total of 152 three-year-olds participated in Experiment 1; 76 children were randomly assigned to each of two conditions: the standard-design condition (mean age = 41.9 months; 36 female) and the relational-design condition (mean age = 41.6 months; 37 female). In each condition, half of the children observed evidence consistent with objects of the same relation producing an effect, and half observed evidence consistent with objects of a different relation producing an effect. In each condition, half of the children observed evidence consistent with objects of the same shape and color. In different training trials, music played only when the objects in a pair were different shapes and colors. In test trials, children were presented with two pairs of novel blocks: one in which the objects were the same and one in which the objects were different. The experimenter recorded whether children selected the pair in which the objects had the same relation (same vs. different) as the objects in the pairs that had caused music to be played during the training trials.

**Fig. 1.** Schematic illustrating the design of training and test trials in Experiments 1 and 2. During training, the experimenter presented pairs of blocks to children; some of the blocks caused music to play when placed on top of a machine (standard-design condition) or into openings on the machine (relational-design condition; see Fig. 2). In same training trials, music played only when the objects in a pair were the same shape and color. In different training trials, music played only when the objects in a pair were different shapes and colors. In test trials, children were presented with two pairs of novel blocks: one in which the objects were the same and one in which the objects were different. The experimenter recorded whether children selected the pair in which the objects had the same relation (same vs. different) as the objects in the pairs that had caused music to be played during the training trials.

**Materials and procedure.** The materials and procedure for the standard-design condition replicated those used by Walker et al. (2016, Experiment 1; see Fig. 1). Children were seated at a table across from the experimenter. The experimenter began by placing an opaque cardboard box on the table, saying “This is my toy! Sometimes when I put things on top, the toy will play music, and other times it does not. Should we try some and see how it works?” As in previous research, the machine appeared to activate and play a novel melody in response to certain combinations of blocks being placed on it. (The machine actually activated a wireless doorbell via a hidden button.) The experimenter was blind to the hypotheses of the study.

Pairs of same and different painted wooden blocks were used during the training trials. After introducing the machine, the experimenter produced two blocks having either the same or a different relation producing an effect. Sample size was predetermined and satisfied a power analysis with power greater than .80, given an alpha of .05 and an effect size (Φ) of .30 (medium). This choice of sample and effect size was based on the findings from previously published data using the identical standard-design task (i.e., the causal-relational-reasoning task) and age group (3-year-olds in the United States; Carstensen et al., 2019), and it is also consistent with other related findings using this method (e.g., Walker et al., 2016). An additional 9 participants were excluded because of experimenter error (n = 3), failure to complete the study (n = 4), parental interference (n = 1), or interference by another child (n = 1). Children were recruited and tested in the lab, at preschools, and at museums. All participants were tested in a quiet, private room with the experimenter.

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Do you hear anything? I don't hear anything." After the second try, she said, "No music. These ones did not make my toy play music." This pattern was repeated with two additional pairs of blocks, one in each relation. The experimenter always began with a causal pair (identical blocks in the same condition and blocks of unique colors and shapes in the different condition) and then alternated inert, causal, and inert pairs, using novel blocks in each new pair. The specific blocks were randomly selected across participants.

After the four training trials, the experimenter said, "Now that you've seen how my toy works, I need your help finding the things that will make it play music. I have two choices for you." The experimenter presented the child with two new pairs composed of novel blocks, one same pair and one different pair. Each pair was presented on a plastic tray, which the experimenter held up, saying, "I have these, and I have these [directing the child's attention to each pair]. Only one of these trays has things that will make my toy play music. Can you point to the tray that has the things that will make it play?" The trays were then placed out of the child's reach on either side of the machine, with each pair set an equal distance from the child. The order and the side of presentation of the correct pair were counterbalanced between participants. The experimenter recorded the child's first point or reach, scoring the response as correct (1) if the child chose the test pair (same or different) that corresponded to his or her training and as incorrect (0) if the child chose the opposite pair.

The materials and procedures for the relational-design condition were identical to those in the standard-design condition, with one critical difference: The design of the machine was modified to include two transparent openings on either side (Fig. 2). The openings were constructed using clear, hard plastic boxes measuring 2 in. x 2 in. During training trials, the experimenter simultaneously inserted pairs of blocks into the two openings (one block on either side) rather than placing them on top of the machine, causing the toy to activate and play music. This was the only difference between the two conditions.

**Results**

The number of correct responses given in each condition is reported in Table 1. Replicating previous work (Walker et al., 2016), our results showed that 3-year-old children in the standard-design condition performed equally poorly on both same trials (53%, 95% confidence interval, or CI = [37%, 69%]) and different trials (40%, 95% CI = [24%, 55%]), p = .357 (two-tailed, Fisher's exact test), odds ratio (OR) = 0.587, 95% CI = [0.236, 1.46]. Overall, children in this condition responded at chance (46%, 95% CI = [35%, 58%]), p = .567 (two-tailed, exact binomial), OR = 0.852. Three-year-olds in the relational-design condition also performed no differently on same and different trials (p > .999, Fisher's exact test), OR = 1.00, 95% CI = [0.388, 2.58]; critically, however, these children succeeded in selecting the test
pair that was consistent with their training (66%, 95% CI = [54%, 76%]), \( p = .008 \) (exact binomial), OR = 1.94. Comparing performance across conditions, we found that children in the relational-design condition significantly outperformed those in the standard-design condition (\( p = .022 \), Fisher’s exact test), OR = 0.446, 95% CI = [0.219, 0.897], in inferring the correct relations.

**Experiment 2**

Given Experiment 1’s finding that children have an early sensitivity to an object’s design, we next examined whether the effects of the same manipulation would impact more entrenched beliefs in adults. Adults perform at ceiling on the standard same–different task used in Experiment 1.\(^2\) We therefore selected a different causal-reasoning task that adults are known to fail. Specifically, adults fail to infer the conjunctive relation (i.e., that two objects together cause a machine to activate) and instead favor the more typical disjunctive relation (i.e., that only one object is needed), despite evidence for the conjunctive rule (Lucas & Griffiths, 2010). Interestingly, preschoolers—who have weaker prior commitments—are better able to infer conjunctive rules from the evidence, outperforming adults on this task (Lucas, Bridgers, Griffiths, & Gopnik, 2014; see also Gopnik, Griffiths, & Lucas, 2015, and Gopnik et al., 2017). These prior results suggest that evaluation of evidence alone may be insufficient for discovery of this relation. In order to facilitate this inference, an object’s design must do more than provide additional data—it must help learners generate a hypothesis they might not have considered otherwise.

If an object’s design were sufficient to influence the hypotheses that adults entertain, we predicted that participants would infer this rule in the relational-design condition. Of course, an alternative explanation for our predicted result would be that, regardless of the evidence, seeing two openings encourages adults to test two blocks. To control for this, we also included a condition in which participants observed evidence that statistically favors the disjunctive rule when the relational-design machine was used. We expected adults in this relational-design control condition to infer that only one object is needed, thus allowing us to rule out the alternative explanation.

### Table 1. Number of Correct and Incorrect Responses in Experiment 1

<table>
<thead>
<tr>
<th>Response</th>
<th>Same trials</th>
<th>Different trials</th>
<th>Total</th>
</tr>
</thead>
<tbody>
<tr>
<td>Correct</td>
<td>25</td>
<td>25</td>
<td>50</td>
</tr>
<tr>
<td>Incorrect</td>
<td>13</td>
<td>13</td>
<td>26</td>
</tr>
<tr>
<td>Correct</td>
<td>20</td>
<td>15</td>
<td>35</td>
</tr>
<tr>
<td>Incorrect</td>
<td>18</td>
<td>23</td>
<td>41</td>
</tr>
</tbody>
</table>

**Method**

**Participants.** A total of 102 adults (mean age = 21.3 years, \( SD = 2.3 \); 77 females) participated in Experiment 2; 34 individuals were randomly assigned to each of three conditions: the standard-design (conjunctive-rule) condition, the relational-design (conjunctive-rule) condition, and the relational-design (disjunctive-rule) control condition. The target sample size was predetermined on the basis of effect sizes from prior work (Lucas et al., 2014: condition \( n = 38 \) in Experiment 1 and 28 in Experiment 2). Adults were recruited from a pool of undergraduate psychology majors at a public university in California and were given course credit in exchange for their participation. An additional 7 participants were excluded because of experimental error (\( n = 6 \)) or failure to attend to the experimental procedures (\( n = 1 \)).

**Materials and procedure.** Except for our introduction of the relation-design machine, the materials and procedure for all conditions were based on those used by Lucas et al. (2014). Materials included nine uniquely shaped purple wooden blocks (triangle, cylinder, cube, sphere, star, pentagon, heart, T shape, semicircle). The blocks were broken into sets of three (e.g., triangle, cylinder, cube); two sets were used during training trials, and one set was used during the ambiguous test. As in Experiment 1, the standard machine was used in the standard-design condition, in which blocks were placed on top of the machine to activate it, and the relational machine was used in the two relational-design conditions, in which blocks were inserted into the transparent openings on the sides of the machine to activate it (Fig. 2). For the familiarization phase, one additional standard-design machine was used that was a different color (pink) and shape (cylinder) than those used in Experiment 1. All machines were activated by a wireless doorbell controlled by the experimenter. The experimenter was blind to the hypotheses of the study.

**Familiarization task.** We began with a warm-up task that included a backward-blocking paradigm (Gopnik & Sobel, 2000) to introduce participants to a causal procedure in which object labels could be ambiguous. Participants in both conditions were shown a machine (standard design) and were told,
We are going to try to figure out which things are wugs. You can't tell that something is a wug just by looking at it, but only wugs have wugness inside of them. Luckily, I brought my wugness machine. The way that this machine works is it turns on and plays music when there is wugness to be detected.

Participants were then shown two objects, which were placed on top of the machine simultaneously. The machine activated, playing music. Both objects were then removed, and one of them was placed back on top of the machine. Again, the machine activated. This provided evidence that one object was a “wug,” and also served to explain away the initial activation, leaving the “wugness” of the second object ambiguous (e.g., see Griffiths, Sobel, Tenenbaum, & Gopnik, 2011). Participants were then asked whether each object was a “wug.” Afterward, the familiarization machine and all materials were removed, and participants proceeded to the main experimental procedure.

Training events. Participants were then introduced to a new machine (standard design or relational design, depending on condition) differing in shape and color from the familiarization machine. The experimenter explained,

Now, we are going to try to figure out which things are blickets. You can't tell that something is a blicket just by looking at it, but only blickets have blicketness inside them. I brought my blicketness machine. The way that the blicketness machine works is it turns on and plays music when there is blicketness to be detected.

Prior to observing objects on the machine, participants were shown a single object and asked whether or not it was a “blicket.” Specifically, they were asked, “First, before we try any objects in my machine, do you think this is a blicket?” This allowed us to empirically estimate how likely participants were to judge that a particular object was a “blicket” when no evidence was available and to ensure that their subsequent responses were significantly different from baseline probability that objects were “blickets.”

Participants were then shown three objects, and they observed a set of training events in which the experimenter placed objects either alone or in pairs on (or in) the machine. In the standard-design condition, objects were placed on top of the machine. In the relational-design conditions, objects were instead inserted into the openings on either side of the machine, as in Experiment 1.3 In some cases, the objects would cause the machine to activate (play music).

The training events followed directly from the conjunctive or disjunctive training used by Lucas et al. (2014); see Figure 3. Depending on condition, participants either observed evidence that two blocks together caused the machine to activate (conjunctive rule) or observed evidence that only a single block was needed (disjunctive rule). Identities of the individual objects that activated the machine and the order of the sets were counterbalanced.
Ambiguous test events. For the ambiguous events, participants were introduced to three new blocks (D, E, and F) and observed the events illustrated in Figure 3 (identical to the procedure of Lucas et al., 2014). Critically, if viewed without the initial training, the events were designed to be ambiguous between a disjunctive (F) or a conjunctive (D and F) causal rule. However, if participants learned the relationship required to activate the “blicketness” machine during the training trials, they could then apply this knowledge to disambiguate the evidence and draw the appropriate inference about the causal status of each of the blocks. That is, we expected individuals in the conjunctive-rule conditions to apply the label to both objects D and F, but not to object E, because both objects D and F would be required to have “blicketness” inside to jointly activate the machine. In contrast, individuals in the disjunctive (relational-control) condition should apply the label only to object F because if only one block is necessary to activate the machine, the evidence of the initial trials rules out D as a possible cause. In all conditions, participants could explain away the E block’s association with activation (on the sixth ambiguous event) by virtue of causally identifying the other blocks from the other trials. Finally, after observing the ambiguous events, participants were asked to infer whether or not each of the new objects was a “blicket.” This served as our critical test measure.

Results

Familiarization and baseline questions. The majority of participants made the backward-blocking inference during familiarization (89%, 95% CI = [82%, 95%]), p < .001 (exact binomial), OR = 8.09; there was no difference across conditions (p = .616, n.s., Fisher’s exact test), suggesting that participants had no trouble reasoning causally about the machines and were able to use information to screen out ambiguous blocks (as was required for block E in the primary experimental task). Second, in the primary experimental task, only 24% of participants inferred that the baseline object was a “blicket,” indicating that adults assumed “blickets” to be relatively rare when compared with chance (50%), (p < .001, exact binomial), 95% CI = [16%, 33%], OR = 0.316; there were no differences across conditions (p = .956, n.s., Fisher’s exact test). This baseline measure also served as a point of comparison for assessing participant responses in the test condition.

Effects of design manipulation on judgments. Results appear in Figure 4, and the exact number of adults who labeled each object a “blicket” is reported in Table 2. There was no difference in participants’ tendency to label object E a “blicket” among the three conditions—standard design (conjunctive rule): M = 6%, 95% CI = [0%, 14%]; relational design (conjunctive rule): M = 12%, 95% CI = [1%, 23%]; relational design (disjunctive rule): M = 9%, 95% CI = [0%, 18%], p = .771 (Fisher’s exact test). Also, as expected, participants in both the standard-design (conjunctive-rule) and relational-design (conjunctive-rule) conditions labeled object F a “blicket” more often than would be expected by chance—standard design (conjunctive rule): M = 79%, 95% CI = [66%, 93%], p < .001, OR = 11.9; relational design (conjunctive rule): M = 88%, 95% CI = [77%, 99%], p < .001, OR = 23.2. There was no difference between groups, p = .512 (OR = 1.94, 95% CI = [0.512, 7.38]).

Critically, however, there was a significant difference in the tendency to label object D a “blicket” across conditions, p < .001 (Fisher’s exact test). Planned comparisons
Table 2. Number of Participants Who Labeled Each Object a “Blicket” in Experiment 2

<table>
<thead>
<tr>
<th>Object</th>
<th>Standard design (conjunctive rule)</th>
<th>Relational design (conjunctive rule)</th>
<th>Relational design (disjunctive rule)</th>
</tr>
</thead>
<tbody>
<tr>
<td>D</td>
<td>6</td>
<td>15</td>
<td>0</td>
</tr>
<tr>
<td>E</td>
<td>2</td>
<td>4</td>
<td>3</td>
</tr>
<tr>
<td>F</td>
<td>27</td>
<td>30</td>
<td>34</td>
</tr>
</tbody>
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revealed that although participants in the standard-design (conjunctive-rule) condition rarely labeled object D a “blicket” ($M = 18\%$, 95\% CI = [5\%, 30\%]), a result not different from chance (binomial, $p = .546$, $OR = 0.695$), those in the relational-design (conjunctive-rule) condition did label object D a “blicket” ($M = 44\%$, 95\% CI = [27\%, 61\%]) more often than would be expected by chance (binomial, $p = .014$, $OR = 2.49$); there was a significant difference between these groups, $p = .034$ (Fisher’s exact test), $OR = 3.68$, 95\% CI = [1.21, 11.2]. In line with our predictions, these findings suggest that the design of the relational machine served to support adult inferences of the conjunctive form. Replicating what would be expected for disjunctive-rule inference, results showed that participants in the relational-design (disjunctive-rule) condition never labeled object D a “blicket” ($M = 0$, 95\% CI = [0\%, 0\%], $p < .001$, binomial), which is significantly different from the finding in the relational-design (conjunctive-rule) condition reported above, $p < .001$ (Fisher’s exact test), $\varphi = -.54$. This demonstrates that the relational-design machine did not simply increase adults’ tendency to label more than one object a “blicket” and therefore cannot explain the improved performance in the relational-design (conjunctive-rule) condition.

General Discussion

Our results demonstrate that both children and adults are sensitive to the design of the learning context when reasoning about causal relationships. Although 3-year-olds in the standard-design condition failed to recognize the relational hypothesis (a result replicating prior work), increasing the salience of this hypothesis through the application of a subtle design cue heightened their tendency to engage in relational reasoning. In addition to providing evidence for the role of design in constraining causal inference, these data provide additional support for the proposal that children’s previous reasoning failures do not result from a lack of competence (e.g., Walker et al., 2016). These results are particularly striking given children’s strong prior for causal hypotheses based on individual objects. In order to use the design of the learning context to override it, these children had to make a sophisticated inference: They had to notice this cue, infer that an object’s design is relevant for its function, and weigh this information more heavily than their prior commitments.

Experiment 2 shows that adults also integrate design features to improve causal inference from evidence. Moreover, our results suggest that adults’ previous failure to infer the conjunctive rule may have been attributable to their failure to consider it as a possible hypothesis during their evaluation of evidence. This is consistent with prior accounts suggesting that evaluation and generation of hypotheses may represent two distinct processes underlying inductive inference (Klahr, Fay, & Dunbar, 1993; Kuhn, 1989). The current experiments provide initial evidence that the design of objects may influence hypothesis generation, changing learning outcomes even for the more entrenched beliefs and biases characteristic of adults. Together, these findings suggest that relatively minor elements of an object’s design can change the distribution of a learner’s prior expectations, constrain the types of hypotheses generated, and influence learning.

Given that object design serves as a constraint for both young children and adults, it is likely a useful cue across the age range. However, future work might examine the nature of these effects in older children and adolescents and might explore a wider range of learning contexts and knowledge domains. Future work might also consider whether similar effects are found in other cultural contexts—particularly those with fewer artifacts. It is possible that the sensitivity to these cues is itself learned through exposure to complex tools and designed environments. On a related note, there are a number of open questions regarding how the particular design modifications used here influenced reasoning. One possibility is that the relational design simply served to disrupt learners’ initial intuitions about the likely causal mechanism, leading them to consider alternatives more broadly. If so, this may have made it more likely for participants to discover the relational hypothesis, albeit indirectly.

Finally, these results have clear implications for the design of learning environments. Our findings dovetail with literature in education that points to the importance of “mise en place” or setting the stage for learning (Weisberg, Hirsh-Pasek, Golinkoff, & McCandliss, 2014). As demonstrated here, both young children and adults are sensitive to an object’s design when engaged in causal inference. These findings therefore open up new avenues for research examining how learning environments can be used to constrain reasoning, support belief revision, and guide discovery.
Transparency

Action Editor: Marc J. Buehner
Editor: D. Stephen Lindsay
Author Contributions
C. M. Walker and E. Bonawitz designed the studies; A. Rett and C. M. Walker performed the research; A. Rett and C. M. Walker analyzed the data; and C. M. Walker, A. Rett, and E. Bonawitz wrote the article. All authors approved the final version of the manuscript for submission.

Declaration of Conflicting Interests
The author(s) declared that there were no conflicts of interest with respect to the authorship or the publication of this article.

Funding
This research was funded by a Hellman’s Fellowship awarded to C. M. Walker.

Open Practices
All data have been made publicly available via the Open Science Framework and can be accessed at https://osf.io/wtptmd/. The design and analysis plans for the experiments were not preregistered. The complete Open Practices Disclosure for this article can be found at http://journals.sagepub.com/doi/suppl/10.1177/0956797619898134. This article has received the badge for Open Data. More information about the Open Practices badges can be found at http://www.psychologicalscience.org/publications/badges.

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Acknowledgments
We are grateful to Mike Frank for his early contributions to the methodology of these experiments. We thank Alicia Lunardhi, Valentina Mancuso, Nicky J. Sullivan, and Emily To for their efforts toward data collection. We also thank the Fleet Science Center, the Birch Aquarium at Scripps Institute of Oceanography, and the New Children’s Museum, as well as participating preschools and families.

Notes
1. This dip in performance is not observed in China (see Carstensen et al., 2019), suggesting that variation in the learning environment leads to differences in relational responding.
2. In a pilot study with 38 adults on Amazon’s Mechanical Turk, 100% selected the correct test pair (same or different) in the standard-design condition.
3. When a training event included a single object, the experimenter placed the object in one side. The choice of side was counterbalanced between sets.
4. When the triplet (D, E, F) was demonstrated with the relational design, the experimenter placed two objects in one opening and one in the other, with object placement counter-balanced between participants.
5. Chance was determined using the baseline measure (24%).
6. There was a significant difference in the tendency to label object F a “blicket” across all three conditions, \( p = .014 \). This was driven by the fact that 100% of adults in the relational-design (disjunctive-rule) condition indicated that object F was a “blicket.”

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