Anxiety disorders and accelerated cellular ageing
Josine E. Verhoeven, Dóra Révész, Patricia van Oppen, Elissa S. Epel, Owen M. Wolkowitz and Brenda W. J. H. Penninx

Background
Anxiety disorders increase the risk of onset of several ageing-related somatic conditions, which might be the consequence of accelerated cellular ageing.

Aims
To examine the association between anxiety status and leukocyte telomere length (LTL) as an indicator of cellular ageing.

Method
Data are from individuals with current (n = 1283) and remitted (n = 459) anxiety disorder, and controls (n = 582) with no psychiatric disorder from the Netherlands Study of Depression and Anxiety. We determined DSM-IV anxiety diagnoses and clinical characteristics by structured psychiatric interviews and self-report questionnaires; LTL was assessed using quantitative polymerase chain reaction and converted into base pairs (bp).

Results
Patients in the current anxiety group (bp = 5431) had significantly shorter LTL compared with the control group (bp = 5506, P = 0.01) and the remitted anxiety group (bp = 5499, P = 0.03) in analyses adjusted for sociodemographics, health and lifestyle. The remitted anxiety group did not differ from the control group (P = 0.84), however, time since remission was positively related with LTL. Furthermore, anxiety severity scores were associated with LTL in the whole sample, in line with a dose–response association.

Conclusions
Patients with current – but not remitted – anxiety disorder had shorter telomere length, suggesting a process of accelerated cellular ageing, which in part may be reversible after remission.

Declaration of interest
E.E. is a cofounder of Telomere Diagnostics, Inc., a telomere measurement company.

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The NESDA sample consists of 2981 participants between 18 and 65 years including people with a current or remitted diagnosis of a depressive and/or anxiety disorder (74%) and healthy controls (26%). To represent various settings and stages of psychopathology, participants with depression and anxiety disorders were recruited at three different locations in The Netherlands in different settings: the community, primary care and specialised mental healthcare settings. People with insufficient command of the Dutch language or a primary clinical diagnosis of other severe psychiatric conditions, such as bipolar disorder, obsessive–compulsive disorder (OCD), post-traumatic stress disorder (PTSD), severe substance use disorder or a psychotic disorder, as reported by themselves or their mental health practitioner, were excluded. Participants were recruited between September 2004 and February 2007. The study was approved by the ethical review board of participating centres, and all participants signed informed consent. Participants in NESDA were assessed during a 4 h clinic visit. The population and methods of the NESDA study have been described in more detail elsewhere.29 For the current study, three groups were created: a control group, a group of participants with remitted anxiety disorders and a group of participants with current anxiety disorders including panic disorder with and without agoraphobia, social phobia, GAD and agoraphobia, thereby excluding 657 participants who did not meet the criteria of one of the three groups. A total of 37 participants were subsequently excluded from analyses because of missing telomere length data, leaving 2324 individuals. Individuals in the control group (n = 582) were defined as having no lifetime history of anxiety or depressive disorders as assessed by the DSM-IV Composite International Diagnostic Interview (CIDI, version 2.1),30 and an anxiety severity score below 11 on Beck’s Anxiety Inventory (BAI).31 Participants in the current anxiety disorder group (n = 1283) had one or more CIDI-diagnosed anxiety disorders (747 had one anxiety diagnosis, 404 had two diagnoses and 132 had three comorbid anxiety disorders) in the past 6 months. The participants in the remitted anxiety disorder group (n = 459) had an anxiety disorder previously but not in the past 6 months according to the CIDI.

**Measurements**

**Telomere length**

Fasting blood was drawn from participants in the morning between 08.30 h and 09.30 h and blood samples were stored in a −80 °C freezer afterwards. Leukocyte telomere length (LTL) was determined at the laboratory of Telomere Diagnostics, Inc. (Menlo Park, California, USA), using quantitative polymerase chain reaction (qPCR), adapted from the published original method by Cawthon et al.32 Telomere sequence copy number in each patient’s sample (T) was compared with a single-copy gene copy number (S), relative to a reference sample. The resulting T/S ratio is proportional to mean LTL.32,33 The detailed method is described elsewhere.34

To compare T/S ratios to telomere restriction fragments (TRF) reported by studies using Southern blot analysis, we used the following steps to derive a conversion formula. Lin et al.34 used a formula of base pairs (bp) = 3274 + 2413 × T/S based on comparison of T/S ratios and TRF analysis of a series of genomic DNA samples from the human fibroblast cell line IMR90. Comparison of the T/S ratios of eight quality control DNA samples from the Telome Health Inc. laboratory that were included on each PCR run, generated the following formula: T/S_Lin et al.34 = (T/S_TelomereDiagnostics) − 0.0545)/1.16. Therefore, the final formula we used to convert T/S ratios to bp is: bp = 3274 + 2413 ((T/S_TelomereDiagnostics) − 0.0545)/1.16. The reliability of the assay was adequate: eight included quality control DNA samples on each PCR run illustrated a small intra-assay coefficient of variation of 5.1% and the interassay coefficient of variation was also sufficiently low (4.6%).

**Psychiatric characteristics**

Four severity measures were included, corresponding with the anxiety disorders in this study: (a) panic disorder arousal symptoms were assessed by the 21-item BAI;31 (b) social phobic symptoms were measured with the five-item social phobia subscale of the Fear Questionnaire,35 (c) the five-item agoraphobia subscale of the Fear Questionnaire assessed agoraphobic symptoms; and (d) GAD’s worrying and rumination were assessed with the Penn State Worry Questionnaire (PSWQ).36 Anxiety duration in recent years was assessed by the Life Chart interview (LCI),37 which uses a calendar method to assess the number of months in which anxious symptoms and avoidance were present during the past 4 years. The CIDI also assessed comorbid major depressive disorder (MDD), which was considered present when diagnosed in the past 6 months. Current use of psychoactive medication was measured by inspection of the active component or medication brand, the dose and frequency of intake from the label on the containers brought in and was categorised using the World Health Organization Anatomical Therapeutic Chemical (ATC) classification38 into antidepressants (tricyclic antidepressants (N06AA), selective serotonin reuptake inhibitors (N06AB) and other antidepressants (N06AF, N06AG, N06AX)) and benzodiazepines (N03Ab, N05BA, N05CD, N05CF).

**Covariates**

Gender, age and years of education were assessed during the interview. Body mass index (BMI) was calculated as measured weight divided by height-squared and divided into underweight (<18.5), normal (18.5–24.9), overweight (25.0–30.0) and obese (>30.0). Alcohol consumption was categorised as non-drinker, moderate drinker (female <14 and male <21 drinks/week) or heavy drinker (female ≥14 and male ≥21 drinks/week). Smoking status was categorised into current, former or never smoker. Physical activity was assessed using the International Physical Activity Questionnaire (IPAQ)39 and expressed as overall energy expenditure in metabolic equivalent total (MET)-minutes per week (MET-level × minutes activity × events per week, see Ainsworth et al.40). The number of current somatic diseases for which participants received medical treatment (i.e. heart disease, epilepsy, diabetes, osteoarthritis, stroke, cancer, chronic lung disease, thyroid disease, liver disease, intestinal disorders and ulcers) was counted.

**Statistical analyses**

Baseline characteristics were compared across anxiety disorder status (control, remitted and current anxiety disorder groups) using chi-squared statistics and analyses of variance (ANOVA). We used analyses of covariance (ANCOVA) to determine differences in LTL in the remitted and current anxiety disorder groups compared with each other and with the control group, controlling for all covariates. Within the remitted anxiety group we tested whether the time since remission (in years) was associated with LTL to test the hypothesis that LTL restores with a longer time since remission. Also, sensitivity analyses were performed in which participants with current and lifetime MDD were excluded to check whether a possible association is present in people with anxiety disorder only and not caused by comorbid MDD. The extra analyses were all fully adjusted. To examine the
role of type of anxiety disorder, adjusted mean LTL of the control group was compared with each of the five disorder subgroups: panic disorder with agoraphobia, panic disorder without agoraphobia, GAD, social phobia and agoraphobia. Cohen’s $d$ was determined as an effect size estimation for significant results. Finally, covariate-adjusted multiple linear regression analyses were used to analyse the association of anxiety characteristics with LTL both within the total sample and within the current anxiety sample. All analyses were conducted using SPSS version 20 for Windows.

Table 1 shows baseline characteristics across anxiety status. The mean age of the total study sample ($n = 2324$) was 41.7 years (s.d. = 13.1, range 18–65) and 66.7% were female. The remitted anxiety disorder group was slightly older than the two other groups, while the control group contained fewer women than both anxiety groups. The current anxiety disorder group were more often less educated, current smokers and non-drinkers or heavy drinkers, and both remitted and current anxiety disorder groups had more somatic diseases than controls. Table 2 shows that the three groups differed on all psychiatric characteristics with the current anxiety disorder group showing the most severe characteristics. These data demonstrate that people with anxiety disorders are in several ways different from non-psychiatric controls and suggest that this epidemiological study sample consists of people with anxiety disorder as actually seen in the community, but that controlling for these differences is insuperable.

LTL was normally distributed and on average 5463 bp (s.d. = 614). It exhibited a significant negative correlation with...
age \( r = -0.326, P < 0.001 \) that corresponded to a mean shortening rate of 14 bp/year. This shortening rate is comparable with previously reported rates based on cross-sectional data (14 bp/year by Cawthon et al., \(^{20}\) 20 bp/year by Hartmann et al.\(^{22}\) and 15–17 bp/year by Wikgren et al.\(^{23}\)). Female participants had longer LTL than male participants (\( F = 14.23; P < 0.001, \) corrected for age). LTL was, besides age and gender, associated with weight (shorter LTL for being underweight, overweight or obese), smoking and drinking status (former and current smokers as well as heavy drinkers have shorter LTL) and number of somatic diseases (more diseases, shorter LTL), but not with education and physical activity.

Between-group analyses showed that the LTL of the control, remitted and current anxiety groups differed in unadjusted \( (F(2,2321) = 3.57, P = 0.028) \), sociodemographic-adjusted \( (F(2,2318) = 5.40, P = 0.005) \) and fully adjusted \( (\text{sociodemographic, health and lifestyle) analyses (} F(2,2313) = 4.22, P = 0.015\)) Post hoc analyses showed that those in the current anxiety group had shorter LTL than healthy controls in unadjusted analyses (Table 3). Further, when adjusted for age, gender and education, LTL was significantly shorter for the current anxiety group \( (bp = 5427) \) compared with the control \( (mean \ bp = 5514, P = 0.003) \) and the remitted anxiety group \( (bp = 5499; P = 0.025) \). Differences remained significant in analyses fully adjusted for sociodemographic, health and lifestyle variables: the current anxiety group had on average 75 bp shorter LTL compared with the control \( (P = 0.01, \) Cohen’s \( d = 0.13) \) and 68 bp shorter LTL \( (P = 0.03, \) Cohen’s \( d = 0.12) \) compared with the remitted group (see Table 3 for the 95% confidence intervals). The remitted anxiety group did not significantly differ from the control group \( (P = 0.84) \). However, within the remitted anxiety group, the time since remission was positively associated with LTL \( (\beta = 0.107, P = 0.024, \) adjusted for all covariates). In further analysis we performed a median-split, resulting in a subgroup that remitted between 6 months to 9 years ago \( (n = 225) \) and a group that were remitted for 10 years or longer \( (n = 220) \). The first subgroup \( (bp = 5372) \) had significantly shorter LTL than the 10+ years subgroup \( (bp = 5529; P = 0.022; \) Cohen’s \( d = 0.22) \), suggesting that the cellular ageing process is in part reversible.

In contrast, in an earlier study in the same cohort we found that individuals who had had depression earlier in life but did not currently meet the MDD diagnosis had shorter telomeres than controls.\(^{24}\) Both the remitted patient groups differed from each other in terms of subthreshold symptomatology and length of remission: remitted patients with MDD scored 18.0 (s.d. = 10.2) on the Inventory of Depressive Symptoms,\(^{24}\) which is still indicative of ‘mild symptomatology’\(^{41}\), whereas the remitted patients with anxiety disorder scored 8.9 (s.d. = 7.3) on BAI, which is below the clinically relevant cut-off score of 11.\(^{42}\) Moreover, the average time to remission in the remitted patients with MDD was 5.2 years (s.d. = 6.2), whereas it was much longer in the remitted anxiety group: 18.0 years (s.d. = 18.7) years.

To test whether the associations found in the present study were not caused by comorbid depressive conditions, sensitivity analyses were performed in which participants with current MDD were excluded. These analyses showed that the 564 participants with current anxiety disorders (but no current MDD) still had significantly shorter LTL than controls \( (P = 0.021; \) Cohen’s \( d = 0.14) \) and marginally significantly shorter telomeres than the participants with remitted anxiety disorders \( (P = 0.078; \) both adjusted for all covariates), indicating that the association is truly present in those with anxiety disorders and not just caused by comorbid depressive conditions. Even when excluding lifetime MDD diagnoses, 268 participants with current anxiety disorders had shorter LTL compared with the controls \( (P = 0.024) \), but not compared with those participants in the remitted anxiety group \( (P = 0.273) \).

To examine the role of type of disorder, mean LTLs for the different anxiety disorder subgroups were each compared with the control group (Fig. 1). To reflect high comorbidity rates, patients with comorbid disorders were included in each anxiety disorder subgroup for which they met the diagnosis. Patients with panic disorder with agoraphobia \( (bp = 5428, P = 0.03, \) Cohen’s \( d = 0.14) \), social phobia \( (bp = 5443, P = 0.04, \) Cohen’s \( d = 0.12) \) and GAD \( (bp = 5404, P = 0.002, \) Cohen’s \( d = 0.18) \) had significantly shorter LTL than the control group. Patients with agoraphobia \( (bp = 5497, P = 0.91) \) and panic disorder without agoraphobia \( (bp = 5458, P = 0.19) \), however, did not differ from controls.

Subsequently, the association between several anxiety characteristics and LTL was examined in the total sample as well as within the 1283 participants in the current anxiety disorder group (Table 4). Within the total sample, severity of anxiety arousal symptoms \( (\beta = -0.056, P = 0.007) \), social phobic symptoms \( (\beta = -0.043, P = 0.03) \) and worrying \( (\beta = -0.041, P = 0.04) \) were associated with shorter LTL. A higher number of current anxiety disorder diagnoses was also significantly associated with shorter LTL in the total sample \( (\beta = 0.047, P = 0.020) \), confirming the difference between people with and without a current anxiety disorder. No LTL-associations were found for symptom duration in the past 4 years \( (P = 0.14) \), comorbid MDD \( (P = 0.54) \) or the use of psychopharmacological medications (Table 4). Also, no anxiety characteristics were associated with LTL within the current anxiety disorder group by itself, suggesting that differences in LTL were only apparent when using the entire spectrum of participants in the healthy control and anxiety groups, but did not further differentiate within the current anxiety group.

**Discussion**

**Main findings**

This large-scale study demonstrated that patients with a current anxiety disorder on average had shorter LTL than non-psychiatric controls and patients with remitted anxiety disorder. Although

<table>
<thead>
<tr>
<th>Table 3 Mean leukocyte telomere length (LTL) in base pairs by anxiety disorder status in unadjusted, basic and fully adjusted analyses</th>
</tr>
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<tbody>
<tr>
<td><strong>LTL, mean (s.e.)</strong></td>
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<tr>
<td>------------------------</td>
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<td></td>
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<tr>
<td><strong>Unadjusted</strong></td>
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<tr>
<td><strong>Basic adjustment</strong></td>
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<tr>
<td><strong>Full adjustment</strong></td>
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<tr>
<td><strong>Remitted anxiety disorder v. control group</strong></td>
</tr>
<tr>
<td>95% CI</td>
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<tr>
<td>-32.4 to 117.8</td>
</tr>
<tr>
<td>22.3 to 117.8</td>
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<tr>
<td>-86.9 to 55.9</td>
</tr>
<tr>
<td>9.1 to 133.6</td>
</tr>
</tbody>
</table>

a. Adjusted for age, gender, education.
b. Adjusted for age, gender, education, body mass index, smoking, alcohol, physical activity and somatic diseases.
Effect sizes were modest, the difference may indicate 3–5 years of accelerated ageing for the current anxiety group, based on the estimated mean telomere shortening rate of 14–20 bp/year as found in our and other studies.20,22,23 The observed associations remained significant after controlling for lifestyle and health variables suggesting that the shorter telomeres were not simply because of an unhealthy lifestyle or physical illnesses within the anxiety group.

Within the current anxiety disorder group, we showed that type of anxiety disorder further differentiated across disorders: patients with panic disorder with agoraphobia, social phobia and GAD, but not panic disorder without agoraphobia and agoraphobia without panic disorder, had significantly shorter LTL than the control group. The differences in LTL for the anxiety subgroups corresponded with an estimated 3.5–5 years (social phobia), 4–6 years (panic disorder with agoraphobia) and 5.5–8 years (GAD) of accelerated ageing compared with the control group. Based on descriptive information (data not shown) it appeared that the two anxiety disorders that did not differ from controls in terms of LTL, showed less severe characteristics compared with other anxiety disorder subgroups. For instance, participants with agoraphobia and panic disorder without agoraphobia tended to have a smaller number of comorbid anxiety disorders, less often comorbid depression and/or slightly lower anxiety severity scores than the other three anxiety subgroups. Consequently, the fact that participants with agoraphobia and panic disorder without agoraphobia reported generally milder symptomatology might explain why LTL was less

Table 4  Associations between leukocyte telomere length and anxiety characteristics in different samples a

<table>
<thead>
<tr>
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<th>Total sample (n = 2324)</th>
<th>Current anxiety disorder group (n = 1283)</th>
</tr>
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<tbody>
<tr>
<td></td>
<td>β</td>
<td>P</td>
</tr>
<tr>
<td>Anxiety severity</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Anxiety arousal (Beck's Anxiety Inventory)</td>
<td>−0.056</td>
<td>&lt;0.01</td>
</tr>
<tr>
<td>Social fear (subscale Fear Questionnaire)</td>
<td>−0.043</td>
<td>0.03</td>
</tr>
<tr>
<td>Agoraphobia (subscale Fear Questionnaire)</td>
<td>−0.021</td>
<td>0.31</td>
</tr>
<tr>
<td>Worrying (Penn State Worry Questionnaire)</td>
<td>−0.041</td>
<td>0.04</td>
</tr>
<tr>
<td>Duration (months with anxiety symptoms in past 4 years)</td>
<td>−0.029</td>
<td>0.14</td>
</tr>
<tr>
<td>Number of current anxiety disorders</td>
<td>−0.047</td>
<td>0.02</td>
</tr>
<tr>
<td>Comorbid major depressive disorder</td>
<td>−0.012</td>
<td>0.54</td>
</tr>
<tr>
<td>Antidepressant use</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Tricyclic antidepressant</td>
<td>−0.006</td>
<td>0.76</td>
</tr>
<tr>
<td>Selective serotonin reuptake inhibitor</td>
<td>−0.033</td>
<td>0.10</td>
</tr>
<tr>
<td>Other antidepressant</td>
<td>−0.036</td>
<td>0.07</td>
</tr>
<tr>
<td>Benzodiazepine use</td>
<td>−0.007</td>
<td>0.73</td>
</tr>
</tbody>
</table>

a. All adjusted for age, gender, education, body mass index, smoking, alcohol, physical activity and somatic diseases.
significantly decreased in these subgroups than in, for example, patients with panic disorder with agoraphobia, social phobia and GAD.

Our results suggest that telomere shortening might in part be reversible since the LTL difference with controls was not present in the group with remitted anxiety disorder. Furthermore, the time since remission was positively associated with LTL within the remitted anxiety group and participants with a remission time under 10 years had shorter telomeres than those who were remitted for 10 years or longer. An earlier study in the same cohort showed that patients with remitted MDD had shorter LTL than controls, suggesting that depression does leave a ‘biological scar’.24 But alternatively, the fact that there was a difference in LTL between the patients with remitted MDD compared with controls, whereas it did not differ in the patients with remitted anxiety disorder, might have been the result of the relatively higher levels of subthreshold symptoms and shorter remission time found in the remitted MDD group.

We further found that anxiety severity scores (anxiety arousal, social fear and worrying) were negatively associated with LTL in the total study sample. However, these variables were not associated with LTL within the current anxiety group, suggesting a dose–response effect across the entire spectrum of healthy controls and participants with anxiety but no further differentiations within the current anxiety group. In line with this, the number of anxiety disorders was also associated only in the total study sample. An explanation for this is that the overall effect of diagnosis status is stronger than that of a dose–response relationship within the participants with anxiety. Alternatively, the severity and duration measures used in our study may not fully represent the ‘anxiety dose’ component in patients. For instance, the anxiety severity measurements used were reflectors of symptom severity over the past week, but not over a much longer period and may therefore not be the best representation of general anxiety severity in patients. Duration of anxiety symptoms was not associated with LTL in this study. This might be as a result of the fact that generally most patients with anxiety disorder reported rather long symptom duration and therefore the range in this variable may not have been enough to illustrate associations with LTL. Comorbid MDD was not associated with LTL either, suggesting that the effects of anxiety disorder status on LTL were independent of MDD diagnosis, which was moreover confirmed by sensitivity analyses that showed that LTL of participants with anxiety disorder still differed from controls when we excluded patients with current and lifetime MDD. We did not find an association between psychopharmacological medication and LTL, but more specific relationships between duration, type and dose of medication remain to be investigated.

Possible mechanisms

Our results align with the general tendency of the limited research into telomere length and anxiety disorders up until now: previous studies found anxiety disorders to be prospectively related to reduced telomere length,50 and telomere length to be shorter in an older subgroup of participants with anxiety disorder compared with a similar-aged control group.51 The observed shorter LTL in this study might be the consequence of several disturbances found in people with anxiety disorders, such as dysregulations in the major bodily stress systems such as the HPA axis and the immune system.40 Dysregulations of these stress systems could contribute to telomere shortening in patients with anxiety disorders.11,42 In line with this, several in vitro and in vivo studies found increased cortisol,44 oxidative stress45 and proinflammatory cytokines46 to be associated with shorter telomere length. However, the exact biological mechanisms that mediate the relationship between certain anxiety disorders and telomere shortening, as well as the direction of the link, remain to be further explored.

Strengths and limitations

The major strengths of the present study are the large sample size, the inclusion of participants with well-characterised current and remitted anxiety disorder diagnoses as well as healthy controls, the wide age range and the assessment of important covariates such as somatic health and lifestyle variables. However, some limitations of the present study should also be noted. First, the associations between anxiety disorder status and LTL were of rather small effect size. However, our effect sizes (Cohen’s \(d = 0.12–0.22\)) are not very different from those described in recent studies on pathophysiological mechanisms, for example, increased inflammatory markers in patients with anxiety disorder (Cohen’s \(d = 0.15–0.19\)),10 and decreased brain-derived neurotrophic factor (Cohen’s \(d = 0.15–0.23\))47 and increased cortisol (Cohen’s \(d = 0.15–0.25\))48 across MDD status. Second, as stress arousal is inherent in people with anxiety disorder, it is possible that the association between anxiety disorder status and telomere length is – partly – mediated by exposure to stressful events and experienced stress. By controlling for stress, the sample might no longer reflect ecologically valid anxiety disorders. It is, however, important to note that the observed telomere shortening might not be specific to anxiety disorders. Third, the sample studied may not fully represent the population with anxiety disorders at large, since for example PTSD and OCD, which can be comorbid with the anxiety disorders in our sample, were exclusion criteria. Fourth, our study design was cross-sectional, which might undermine the complexity of telomere length regulating mechanisms, and which does not allow one to draw conclusions regarding causality. Even though we looked at patients with current and remitted anxiety disorders, suggestive of a temporal effect, this study is solely a reflection of the current state of different participants. Since longitudinal studies show complex telomere length dynamics with both shortening as well as lengthening of telomeres over time,49–52 future studies should explore the relationship longitudinally. It should also be noted that our variable base pairs is an estimate rather than a directly measured variable. Also, as in nearly all studies, we used leukocytes for telomere length measurement, which is a validated, non-invasive and often-used indicator for cellular ageing. It would, however, be worthwhile to examine cellular ageing processes in other tissues. Finally, telomerase activity was not measured in the current study, but information regarding telomere repair and maintenance would be of great value in future research, since this would give us a more dynamic picture of the telomere/telomerase system in affective disorders.

Clinical implications and further research

This large cohort study provides suggestive evidence that patients with anxiety disorders have shorter telomere length compared with healthy controls, suggesting a pathway of accelerated cellular ageing within the most common anxiety disorders. This might in turn explain the increased somatic health problems found in this population. As participants with remitted anxiety disorder resembled the healthy controls, in terms of telomere length, it is possible that this telomere shortening can be reversed, although the present data did not allow assessment of within-participant changes in telomere length. Lifestyle interventions such as increased physical activity have been shown to have a favourable impact on cellular ageing in a healthy population.53–55
Therefore, further studies should test whether such interventions could result not only in a reversal of psychiatric symptomatology but also in restoring cellular ageing and consequent somatic health.

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First received 20 Dec 2013, final revision 2 May 2014, accepted 15 Aug 2014

Funding

The infrastructure for the USDA study (www.nesda.nl) is funded through the Geestkracht program of the Netherlands Organization for Health Research and Development (ZonMW, grant number 10-000-1002) and is supported by participating universities and mental health care organizations (VU University Medical Center, GGZ in Geest, Arkin, Leiden University Medical Center, GGZ Rivierduinen, University Medical Center Groningen, Leiden, GGZ Friesland, GGZ Drenthe, Institute for Quality of Health Care (IQ healthcare), Netherlands Institute for Health Services Research (NIVEL) and Netherlands Institute of Mental Health and Addiction (Trimbos). B.W.J.H.P., J.E.V., D.R. and telomere length assay were supported through a NWO-VK3 grant (number 91811602).

References

22 Hartmann N, Boehmer M, Groenen F, Kalb R. Telomere length of patients with major depression is shortened but independent from therapy and severity of the disease. Depress Anxiety 2010; 27: 1111–6.
I, me, my cell phone

Shabbir Amanullah

Charlie Brown’s best friend Linus has captured the imagination of millions of fans all over the world. No single professional group has probably found as much pleasure in them as those in the field of mental health. The role of a transitional object is often to provide security, comfort in a world that is very often unpredictable. Even among adults, there are many who still have objects that provide comfort.

The more recent phenomenon of ‘constant technology’ may be seen as essential for most, but is creating a new generation who find it hard to imagine life without a cell phone. When out of range or charge, a number of parents report their teenage children admit feeling anxious. Constant social media updates with the now ubiquitous ‘selfies’ are the norm. Little planning or thought goes into it. It is an impulsive need to respond immediately that is creeping in. Technology has created a new reality that requires an alternative to the ‘blue blanket’.

Cell phones are no longer just tools of communication. They are lifelines for most of us and increasingly, transitional objects for ‘generation cell’. I often wonder, if Linus existed in real life, would he give his blue blanket up for the cell phone? Reality is now arguably in 3D – I, me, mycell!!

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Access the most recent version at DOI: 10.1192/bjp.bp.114.151027

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