Emotion

Trajectories of Affective Well-Being and Survival in Middle-Aged and Older Adults

Dakota W. Cintron and Anthony D. Ong

Online First Publication, January 8, 2024. https://dx.doi.org/10.1037/emo0001341

CITATION

Cintron, D. W., & Ong, A. D. (2024, January 8). Trajectories of Affective Well-Being and Survival in Middle-Aged and Older Adults. Emotion. Advance online publication. https://dx.doi.org/10.1037/emo0001341
Trajectories of Affective Well-Being and Survival in Middle-Aged and Older Adults

Dakota W. Cintron1, 2 and Anthony D. Ong1, 2

1 Department of Psychology, Cornell University
2 Center for Integrative Developmental Science, Cornell University

Affective experiences are key components of subjective well-being with important implications for health. However, little is known about heterogeneous longitudinal affect trajectories and their links to survival. This study identified joint trajectory subgroups based on 18-year changes in positive affect (PA) and negative affect (NA) and examined their differential associations with mortality risk. Participants were 3,250 adults (aged 39–93 years) from the Midlife in the U.S. study assessed over three waves (1995–2013). Parallel growth mixture modeling revealed three subgroups: (a) improving (increasing PA, decreasing NA), (b) deteriorating (decreasing PA, increasing NA), and (c) flourishing (high, stable PA, low, stable NA). Adjusting for baseline demographic and health covariates, Cox proportional-hazard results showed the improving group had the lowest mortality risk (HR = 0.82, 95% CI [0.35, 1.32]) and the deteriorating group had the highest mortality risk (HR = 1.86, 95% CI [1.34, 3.55]), relative to flourishing. These findings highlight the importance of modeling multidimensional trajectories of affective well-being and their heterogeneous links to survival.

Keywords: positive affect, negative affect, survival, growth mixture modeling

Supplemental materials: https://doi.org/10.1037/emo0001341.supp

In the current study, we aimed to identify subgroups exhibiting distinct longitudinal patterns of PA and NA states. Our focus was on trajectories of experienced positive and negative emotions rather than the broader constructs of psychological and social flourishing, encompassing facets like purpose in life, close relationships, and social integration (Keyes, 1998, 2007; Ryff & Singer, 2000). We aimed to elucidate the heterogeneous affective pathways linking positive and negative emotional experiences to mortality risk across adulthood. Identifying joint trajectory subgroups defined by shifts in PA and NA can reveal how these core affective components combine to influence well-being and health outcomes over the life course (Cacioppo & Berntson, 1999). This approach moves beyond examining emotional trajectories in isolation to capture the interconnected nature of PA and NA (Reich et al., 2003).

Several theoretical frameworks provide insights into heterogeneous emotional development across adulthood and the need to examine longitudinal affective patterns. For instance, socioemotional selectivity theory (SST; Carstensen et al., 1999) proposes that as time horizons shrink, adults prioritize emotional fulfillment through positive experiences and meaningful relationships. This motivational shift may manifest as improving affective well-being, with older adults selectively engaging in activities that enhance PA and reduce NA. In contrast, strength and vulnerability integration theory (SAVI; Charles, 2010) posits that aging involves a dynamic interplay between emotional strengths (e.g., regulation, positivity) and vulnerabilities (e.g., losses). This framework predicts heterogeneity in affective trajectories, encompassing stability, improvement, or decline depending on contextual factors. Finally, dynamic integration theory (DIT; Labouvie-Vief, 2003) emphasizes negotiating the balance between complexity and intensity of emotions. DIT proposes that negotiating this balance is a key developmental task, implying potential
fluctuations in both PA and NA over time as integration and differentiation evolve. Together, these theoretical frameworks point to the dynamic nature of affect and the need to examine heterogeneous trajectories, which may show stability, improvement, or variability in emotional well-being.

**Affective Well-Being Trajectories**

Although prior studies have examined PA or NA trajectories, few have modeled their dynamic co-occurrence. For instance, Zaninotto et al. (2016) showed that sustained enjoyment of life over time predicted longer survival, while Willroth et al. (2020) found that increases in PA uniquely predicted better health. While these studies offer insights into the relationships between specific affective components and health, there remains a need for more comprehensive investigations into individual variations in affective well-being trajectories over time. Modeling the dynamic co-occurrence of PA and NA may reveal critical information about their interconnected nature and combined influence on outcomes (Cacioppo & Berntson, 1999; Reich et al., 2003). Identifying subgroups based on distinct joint trajectories can provide targeted insights into the multifaceted affective pathways that shape well-being and health over time (Keyes, 2007; Ryff et al., 1998).

**Joint Modeling of PA and NA**

Person-centered approaches are well-suited for identifying distinct trajectory subgroups (Bussere et al., 2009; Howard & Hoffman, 2018; Laursen & Hoff, 2006; Nagin & Tremblay, 2005). These subgroups reflect variations in trajectories of affective well-being. Some individuals exhibit patterns of maintaining or improving affective well-being over time, while others may exhibit patterns of declining affective well-being. Recognizing these diverse subgroups can inform targeted interventions to promote positive emotional states and mitigate the potential adverse health effects arising from poor affective well-being.

Recent studies reveal subgroups showing stable, increasing, or decreasing well-being trajectories relate differentially to health outcomes (Guimond et al., 2022; Lim et al., 2017; Radler et al., 2018). Notably, these studies focused on psychological and evaluative well-being (e.g., purpose in life, life satisfaction), and it remains unclear whether longitudinal trajectories of affective well-being demonstrate similar protective associations. Focusing on affective well-being trajectories provides insights into the dynamic interplay between PA and NA, enhancing understanding of the multifaceted nature of affective well-being and its implications for overall health. Moreover, unlike previous studies modeling PA and NA separately using growth curve modeling (Willroth et al., 2020) or psychological well-being facets separately using latent class growth modeling (Guimond et al., 2022; Nagin & Tremblay, 2005), we use parallel growth mixture modeling (Muthén & Muthén, 1998–2017; Ram & Grimm, 2009; Wickrama et al., 2021) to examine the combined influence of heterogeneous PA and NA trajectories on survival. This joint modeling approach assesses how combinations of heterogeneous PA and NA trajectories influence survival, elucidating their interconnected nature.

**Affective Well-Being and Mortality**

Affective well-being shows established connections to physical health and survival (Chida & Steptoe, 2008; Pressman & Cohen, 2005). Although some evidence links high, stable PA to lower mortality (Willroth et al., 2020; Zaninotto et al., 2016), heterogeneity in these associations remains underexplored. Examining joint trajectories of PA and NA can elucidate multidimensional pathways linking emotions and survival. For instance, stable high PA and low NA may reflect affective pathways associated with survival benefits, whereas declining PA and rising NA may reflect affective patterns associated with greater mortality risk.

Theoretical explanations for the ties between affective well-being and mortality posit that high PA may benefit physiological functioning through improved immune and cardiovascular systems, while high NA can dysregulate these systems through autonomic, neuroendocrine, and inflammatory pathways (Boehm & Kubzansky, 2012; Pressman & Cohen, 2005). Empirically, high PA has been associated with lower cortisol output, reduced cardiovascular disease risk, and heightened antibody response, whereas high NA is linked to elevated inflammation and poorer cardiovascular health (see Steptoe et al., 2005, 2007). Such effects are believed to cumulatively influence morbidity and mortality over the long term.

**The Current Study**

This study examined the association between latent trajectories of PA and NA and mortality risk in a large sample of U.S. adults. We used measures of PA and NA assessed three times over 18 years and analyzed associations with mortality in a large, national sample of older men and women from the Midlife in the United States (MIDUS) study. Building on prior theoretical frameworks (Carstensen et al., 1999; Charles, 2010; Labouvie-Vief, 2003) and empirical studies (Guimond et al., 2022; Lim et al., 2017; Radler et al., 2018), we defined flourishing trajectories as maintaining relatively high, stable levels of PA and low, stable levels of NA over time. Deteriorating trajectories were conceptualized as exhibiting marked declines in PA and increases in NA over the follow-up period. Finally, improving trajectories were characterized by showing increasing PA and decreasing NA over time.

We hypothesized that adults exhibiting flourishing trajectories would have the greatest survival over the follow-up period, as they may benefit from both the protective effects of PA and the absence of harmful effects of NA. Conversely, those with deteriorating trajectories were expected to show the highest mortality risk, as they may suffer from both the lack of protective effects of PA and the presence of harmful effects of NA. Adults with improving trajectories were expected to show intermediate survival compared to flourishing and deteriorating subgroups, reflecting gradual gains in PA’s beneficial effects and reductions in NA’s adverse influences.

**Method**

**Study Population**

Data for this study were drawn from the MIDUS study. The MIDUS cohort is a national longitudinal sample of U.S. adults (N = 7,108; age = 24–74 years) recruited from 1995 to 1996 (MIDUS 1). Participants from MIDUS 1 were reinvited for follow-up assessments in 2004–2006 (MIDUS 2) and 2013–2015 (MIDUS 3). The current analyses used data from MIDUS participants (N = 3,250; age = 39–93 years) who completed assessments at all three waves and did not have complete missing data on the PA and NA measures. Data collection for the MIDUS studies was approved by the Institutional Review Boards at each participating site, and all participants provided informed consent.
Measures

Affective Well-Being

Participants were asked to rate how much during the past 30 days they felt each of six PA items (cheerful, in good spirits, extremely happy, calm and peaceful, satisfied, and full of life) and six NA items (so sad nothing could cheer you up, nervous, restless, nervous, everything was an effort, and worthless). Item responses were Likert-scaled and ranged from 1 (all of the time) to 5 (none of the time). We reverse-coded PA and NA items such that higher PA scale scores indicated greater PA and higher NA scale scores indicated greater NA. Across the three waves, Cronbach’s αs ranged from .90 to .91. Prior research indicates that these PA and NA measures have evidence of strong psychometric properties, including strong (scalar) invariance across age and gender groups and over time (Chan et al., 2020; Joshanloo & Baksh, 2016; Mroczek & Kolarz, 1998).

Mortality and Survival

Data on mortality were collected using three methods: (a) National Death Index reports conducted in 2006, 2009, and 2016; (b) tracing and mortality closeout interviews conducted by the University of Wisconsin Survey Center as part of MIDUS 3 (2013–2015); and (c) deaths recorded during regular longitudinal sample maintenance time (Chan et al., 2020; Joshanloo & Baksh, 2016; Mroczek & Kolarz, 1998).

Demographic Covariates

Chronological age in years was computed by subtracting birthdates from the baseline interview date. Self-reported sex (male or female), employment (currently working or not), race (non-White or White), and marital status (married or not) were collected at baseline. Income at baseline was assessed from household total income from wage, pension, social security, and other sources. We log-transformed income to account for positive skew. Education at baseline was assessed on a 12-point scale ranging from 1 (no school/some grade school) to 12 (PhD, MD, JD, or other professional degree).

Health Covariates

Self-reported general health was collected at baseline with the question, “Using a scale from 0 to 10 where 0 means ‘the worst possible health’ and 10 means ‘the best possible health,’ how would you rate your health these days?” The number of self-reported chronic health conditions in the past 12 months was collected at baseline. Body mass index (BMI) at baseline was calculated by dividing the respondent’s weight (mass) in kilograms by heights in meters squared. Basic difficulties with activities of daily living (ADL) at baseline were measured using a scale constructed from seven items asking, “How much does your health limit you in doing each of the following?” and include, for example, “lifting or carrying groceries,” “bending, kneeling, or stooping,” and “walking several blocks.”

Statistical Analysis

Analyses proceeded in three steps. To identify multiple trajectory subgroups, we first used latent growth curve modeling (McCoach & Cintron, 2022) to estimate the overall trajectory of PA and NA over 18 years. We fit three unconditional latent growth models for PA and NA (i.e., without covariates): an intercept-only model, a linear model, and a quadratic model. A linear trajectory for both PA and NA was selected based on the Bayesian information criterion (BIC; Schwarz, 1978).

Second, based on results from the first step, we used parallel growth mixture modeling (Muthén & Muthén, 1998–2017; Ram & Grimm, 2009; Wickrama et al., 2021) to identify subgroups of adults with distinct PA and NA joint trajectories. This approach fits a mixture model to the PA and NA growth trajectories where each resultant joint trajectory combines a trajectory for PA and a trajectory for NA (see Figure S1 in the online supplemental materials). To handle missing data in the PA and NA measures during class enumeration, we used a full-information maximum likelihood estimator in Mplus Version 8.8 with robust standard errors under the missing at-random assumption (Enders & Bandalos, 2001; Muthén & Muthén, 1998–2017).

Several indicators were used to select the optimal number of joint trajectory subgroups: (a) entropy index, (b) BIC, (c) sample size adjusted BIC (aBIC), (d) the bootstrapped likelihood ratio test (BLRT), and (e) the Vuong–Lo–Mendell–Rubin (VLMR) likelihood ratio test. Higher entropy values indicate better class separation. Lower BIC and aBIC values indicate better model fit. Significant BLRT or VLMR p values (p < .05) suggest the k-class model fits better than the k–1-class model (Nylund et al., 2007; Tein et al., 2013). As growth mixture modeling is exploratory, our model selection procedures involved balancing statistical fit, parsimony, stability, and substantive utility (Grimm et al., 2016; Weiss et al., 2023; Wickrama et al., 2021). The goal was to identify the optimal number of classes that best described the data while balancing complexity, generalizability, and meaningfulness. Models with higher entropy were preferred as prior research indicates the importance of entropy for correct class assignment, with greater entropy corresponding to less overlap between latent classes and better correct class assignment (Cintron et al., 2023). Likewise, more parsimonious models satisfying the selection criteria were preferred, with priority given to the VLMR and BLRT tests given it is not uncommon for information indices (i.e., aBIC/BIC) to decrease for each additional class added (i.e., there is not a global minimum; Nylund-Gibson & Choi, 2018).

When fitting the models, the NA slope variance was constrained to zero, and factor/residual variances were held equal across classes to aid convergence and obtain admissible solutions, as is common in practice (Diallo et al., 2016; Kooken et al., 2019; McNeish et al., 2023). Overall, the class enumeration process aimed to select the most appropriate model that provided meaningful insights into heterogeneous affective well-being trajectories.

Third, after identifying joint trajectory subgroups, Cox proportional-hazard models (Asparouhov et al., 2006) evaluated the association between group membership and survival time, before and after adjustment for baseline demographics (i.e., age, gender, race, employment status, marital status, education, income) and health (self-rated health, chronic conditions, BMI, ADL difficulties; see Figure S1 in the online supplemental materials). The profile likelihood method in Mplus 8.8 was used for model estimation (Asparouhov et al., 2006). Survival time was right censored for participants who were still living as of December 2021. Bias-adjusted three-step procedures estimated the association between latent class growth trajectories and survival time (Bakk et al., 2013; Vermunt, 2010).
Two sensitivity analyses were conducted. First, parallel censored and censored-inflated growth mixture models evaluated potential floor effects of NA on class enumeration. Second, analyses excluding deaths within 12 and 24 months of the final assay measurement were performed to check for reverse causality (Steptoe, 2019), where declining well-being may reflect approaching end of life. These analyses provided checks to ensure the identified trajectories reflected long-term affective well-being patterns rather than artifacts of floor effects or imminent mortality.

**Transparency and Openness**

We report how sample size was determined, all data exclusions, manipulations, and measures per JARS guidelines (Kazak, 2018). Study data are publicly available online at [https://midus.colectica.org/](https://midus.colectica.org/). Analysis code and materials are available at [https://osf.io/waf2m/?view_only=a32db5fa215433f1a8696664e51e1c0](https://osf.io/waf2m/?view_only=a32db5fa215433f1a8696664e51e1c0) (Cintron & Ong, 2023). The study design and analysis were not preregistered.

**Results**

**Descriptive Statistics**

Table 1 summarizes sample demographics, descriptive statistics for key variables, mortality rate, and average survival time in months. PA and NA scores spanned the full 1–5 range. Of the 3,250 participants, 428 (13%) died over the follow-up period. The average survival time was 101 months. Details on the full versus analytic sample can be found in Table S1 in the online supplemental materials.

<table>
<thead>
<tr>
<th>Variable</th>
<th>Data Count/M (percentage/SD)</th>
</tr>
</thead>
<tbody>
<tr>
<td>N</td>
<td>3,250</td>
</tr>
<tr>
<td>Female</td>
<td>1,792 (55.14)</td>
</tr>
<tr>
<td>Age</td>
<td>45.67 (11.40)</td>
</tr>
<tr>
<td>Non-White</td>
<td>209 (7.00)</td>
</tr>
<tr>
<td>Education</td>
<td>7.31 (2.44)</td>
</tr>
<tr>
<td>Not employed</td>
<td>1,040 (32.13)</td>
</tr>
<tr>
<td>Married</td>
<td>2,357 (72.52)</td>
</tr>
<tr>
<td>Income (log)</td>
<td>10.85 (1.60)</td>
</tr>
<tr>
<td>Self-rated health</td>
<td>7.60 (1.47)</td>
</tr>
<tr>
<td>Number of chronic conditions</td>
<td>2.22 (2.29)</td>
</tr>
<tr>
<td>Body mass index</td>
<td>26.57 (5.11)</td>
</tr>
<tr>
<td>Difficulty with ADL</td>
<td>1.45 (0.64)</td>
</tr>
<tr>
<td>Time 1 positive affect</td>
<td>3.41 (0.70)</td>
</tr>
<tr>
<td>Time 2 positive affect</td>
<td>3.45 (0.69)</td>
</tr>
<tr>
<td>Time 3 positive affect</td>
<td>3.43 (0.72)</td>
</tr>
<tr>
<td>Time 1 negative affect</td>
<td>1.51 (0.58)</td>
</tr>
<tr>
<td>Time 2 negative affect</td>
<td>1.48 (0.54)</td>
</tr>
<tr>
<td>Time 3 negative affect</td>
<td>1.47 (0.58)</td>
</tr>
<tr>
<td>Mortality (deaths)</td>
<td>428 (13.0)</td>
</tr>
<tr>
<td>Survival (months)</td>
<td>101.10 (20.13)</td>
</tr>
</tbody>
</table>

*Note.* All covariates were measured at MIDUS Wave 1 (i.e., baseline). ADL = activities of daily living; MIDUS = Midlife in the United States.

**Parallel Growth Mixture Modeling**

Parallel growth mixture models estimating 1–10 classes were fitted. Model fit information is reported in Table 2. A three-class solution was selected based on the VLMR test, high Entropy, and substantive interpretability of the classes. Figure 1 displays the estimated unconditional three-class growth trajectories: (a) improving (7.7%, n = 214); (b) deteriorating (9.1%, n = 231), and (c) flourishing (83.1%, n = 2,667).

Table 3 (Panel A) and Figure 1 summarize the three-class growth model. The improving subgroup exhibited moderate levels of PA (intercept = 2.40, p < .01) and NA (intercept = 2.84, p < .01) at baseline, with large increases in PA (slope = 0.39, p < .01) and decreases in NA (slope = −0.58, p < .01) over time. The deteriorating subgroup displayed moderate initial levels of PA (intercept = 3.21, p < .01) and NA (intercept = 1.73, p < .01), with marked decreases in PA (slope = −0.34, p < .01) and increases in NA (slope = 0.54, p < .01) over time. Finally, the flourishing subgroup showed high levels of PA (intercept = 3.53, p < .01) and low levels of NA (intercept = 1.36, p < .01) at baseline and relatively stable patterns of PA (slope = 0.01, p < .11) and NA (slope = −0.03, p < .01) over time.

**Survival Analysis**

Hazard ratios (HRs) from the Cox models indicate relative mortality risk between affective well-being trajectories where the flourishing subgroup is the reference group. For example, the HR of 1.65 below signifies that the deteriorating group has approximately a 65% greater risk of experiencing mortality compared to the flourishing group. Confidence intervals provide information on estimate precision and uncertainty. Narrower confidence intervals indicate more precise estimates, while wider intervals reflect greater uncertainty. Table 3 (Panel B) shows Cox regression results. Unadjusted HRs were 0.78 (95% CI [0.52, 1.16]) for improving and 1.65 (95% CI [1.24, 2.18]) for deteriorating subgroups relative to the flourishing subgroup. Adjusting for baseline demographics and health, HRs were 0.82 (95% CI [0.35, 1.32]) and 1.86 (95% CI [1.34, 3.55]), respectively. Although survival rates did not significantly differ between improving and flourishing subgroups, the flourishing and improving subgroups had significantly higher odds of survival compared to the deteriorating subgroups. The unadjusted Kaplan–Meier curves in Figure 2 depict 108-month survival probabilities, with the improving subgroup having the lowest probability of survival over follow-up.

Table S2 in the online supplemental materials displays the effects of demographics and baseline health on latent class membership. Older, married adults in better health were more likely to be in flourishing versus improving subgroups. Additionally, those who were older, more educated, and had higher income were more likely to be in the flourishing rather than the deteriorating subgroup. Income was the only predictor distinguishing improving and deteriorating subgroups, with higher incomes among those in the improving subgroup. Importantly, controlling for differential effects of the covariates did not substantively alter the survival analysis results.

**Sensitivity Analyses**

A sensitivity analysis evaluated potential floor effects on class enumeration by comparing the initial analysis to parallel censored and censored-inflated growth mixture models. The goal was to assess whether the identified subgroups substantially differed across approaches, particularly for NA where floor effects may be present. The improving, deteriorating, and flourishing classes...
were consistently replicated across the initial and censored models (Tables S3–S6 in the online supplemental materials), indicating the robustness of these substantive trajectories. However, the censored models also resulted in extraneous subclasses overextracting the flourishing and improving subgroups (Tables S4 and S6 in the online supplemental materials). Since the key classes were stable, the initial three-class solution was retained for survival analysis, suggesting floor effects did not meaningfully impact enumeration. A second sensitivity analysis omitted deaths within 12 (n = 17) and 24 (n = 73) months of the final assessment to test reverse causation. Associations between joint trajectories and mortality remained unchanged (Tables S7 and S8 in the online supplemental materials), further demonstrating robustness.

### Table 2

<table>
<thead>
<tr>
<th>Classes</th>
<th>Parameters</th>
<th>LL</th>
<th>Entropy</th>
<th>BIC</th>
<th>aBIC</th>
<th>BLRT</th>
<th>VLMR</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>16</td>
<td>-14,187</td>
<td>—</td>
<td>28,543</td>
<td>28,453</td>
<td>—</td>
<td>—</td>
</tr>
<tr>
<td>2</td>
<td>21</td>
<td>-13,625</td>
<td>0.91</td>
<td>27,420</td>
<td>27,353</td>
<td>&lt;.01</td>
<td>&lt;.01</td>
</tr>
<tr>
<td>3</td>
<td>26</td>
<td>-13,092</td>
<td>0.91</td>
<td>26,394</td>
<td>26,312</td>
<td>&lt;.01</td>
<td>&lt;.01</td>
</tr>
<tr>
<td>4</td>
<td>31</td>
<td>-12,910</td>
<td>0.92</td>
<td>26,071</td>
<td>25,972</td>
<td>&lt;.01</td>
<td>0.10</td>
</tr>
<tr>
<td>5</td>
<td>36</td>
<td>-12,759</td>
<td>0.90</td>
<td>25,810</td>
<td>25,695</td>
<td>&lt;.01</td>
<td>0.60</td>
</tr>
<tr>
<td>6</td>
<td>41</td>
<td>-12,622</td>
<td>0.90</td>
<td>25,576</td>
<td>25,446</td>
<td>&lt;.01</td>
<td>0.20</td>
</tr>
<tr>
<td>7</td>
<td>46</td>
<td>-12,518</td>
<td>0.91</td>
<td>25,408</td>
<td>25,262</td>
<td>&lt;.01</td>
<td>0.10</td>
</tr>
<tr>
<td>8</td>
<td>51</td>
<td>-12,445</td>
<td>0.90</td>
<td>25,302</td>
<td>25,140</td>
<td>&lt;.01</td>
<td>0.50</td>
</tr>
<tr>
<td>9</td>
<td>56</td>
<td>-12,367</td>
<td>0.89</td>
<td>25,187</td>
<td>25,009</td>
<td>&lt;.01</td>
<td>0.20</td>
</tr>
<tr>
<td>10</td>
<td>61</td>
<td>-12,290</td>
<td>0.88</td>
<td>25,073</td>
<td>24,879</td>
<td>&lt;.01</td>
<td>0.20</td>
</tr>
</tbody>
</table>

Note. The choice for three classes was based on the VLMR. LL = log-likelihood; BIC = Bayesian information criterion; aBIC = sample size adjusted BIC; BLRT = bootstrapped likelihood ratio test p value; VLMR = Vuong–Lo–Mendell–Rubin adjusted likelihood ratio test p value.

### Figure 1

*Latent Growth Trajectory Classes of Affective Well-Being*

Negative and Positive Trajectories Over Time

- **Improving (7.7%)**
- **Deteriorating (9.1%)**
- **Flourishing (83.1%)**

*Note.* See the online article for the color version of this figure.

---

**Discussion**

Affective states are key components of subjective well-being and play an important role in the maintenance of health. Little is known, however, about the heterogeneous pathways by which affective well-being and survival are linked over the life course. To our knowledge, this is the first study to (a) describe joint trajectory subgroups of PA and NA across adulthood and (b) examine the relationship between joint trajectory subgroups and mortality. The results revealed three primary multidimensional trajectories over 18 years: improving (increasing PA and decreasing NA), deteriorating (declining PA and rising NA), and flourishing (stable high PA and stable low NA).

Our identification of flourishing, deteriorating, and improving affective trajectories provides nuanced insights that align with theoretical perspectives on emotional aging. Consistent with past research on psychological well-being (Guimond et al., 2022; Lim et al., 2017; Radler et al., 2018), the flourishing pattern marked by sustained high PA and low NA was most prevalent. The stability of high PA and low NA in this subgroup reflects continued emotional regulatory strengths associated with lower mortality risk. However, distinct deteriorating and improving subgroups also emerged, elucidating heterogeneous affective pathways. The improving trajectory, characterized by increasing PA and decreasing NA over time, aligns with SST (Carstensen et al., 1999). This motivational shift toward meaningful emotional experiences in later life was associated with reduced mortality risk, suggesting potential health benefits. In contrast, the deteriorating trajectory exhibited marked declines in PA and rises in NA, denoting failures to achieve emotional goals with advancing age. This subgroup had the highest mortality risk, indicating adverse health impacts.

Together, the flourishing, deteriorating, and improving patterns reveal nuanced interplay between diverse affective pathways and
longevity. Our findings provide empirical insights that align with the SAVI model’s premise of heterogeneous affective trajectories differentially linked to survival (Charles, 2010). These trajectories suggest aging involves a dynamic interplay of regulatory strengths, vulnerabilities, and motivational shifts in influencing emotional health.

Critically, while overall affective trajectories show ties to mortality (Willroth et al., 2020; Zaninotto et al., 2016), identifying multiple subgroups informs targeted interventions. Our joint trajectory analysis suggests that adults in the deteriorating subgroup were significantly more likely to die over the follow-up period. For these adults, interventions aiming to promote PA and reduce NA over time may be warranted. By comparison, adults whose trajectories reflected marked increases (improving subgroup) or persistently high (flourishing subgroup) affective well-being may benefit from interventions that sustain these favorable patterns over time (Keyes, 2007; Zaninotto et al., 2016).

Limitations and Future Research

This is an observational study, so issues of selection bias, confounding, or reverse causality may limit our ability to draw causal conclusions. For instance, it is known that there is selection bias in who returns in longitudinal studies. However, the full sample and analytic samples showed comparable demographics, health characteristics, and PA and NA at baseline (see Table S1 in the online supplemental materials). Moreover, we adjusted for several potential confounders, including demographic variables and baseline health to protect against these issues. Since later health issues could explain decreasing PA and increasing NA, reverse causality is possible. However, associations were maintained when excluding deaths.
within 1–2 years of the final affect assessment, arguing against terminal decline effects.

Several modeling constraints should be noted. The three-wave design precluded identifying complex nonlinear patterns, restricting generalizability (Ram & Grimm, 2009). Additional assessment waves could enable more reliable detection of unobserved heterogeneity. Constraining residual variances and fixing the NA slope variance addressed convergence issues but restricted full parameterization (Bauer & Curran, 2003; Diallo et al., 2016). While sensitivity analyses indicated the classes were reasonably robust to floor effects, future research could apply joint covariance pattern mixture models (McNeish et al., 2023). Replication with larger subgroups is important given the small sample sizes in some subgroups (Cintron et al., 2023).

Looking ahead, elucidating contextual influences on affective trajectories represents a critical goal, as does exploring moderators of the affect-health linkage (Carstensen et al., 1999; Charles, 2010). Analyses of interaction effects between PA and NA may offer additional insights into their interconnected nature (Cacioppo & Berntson, 1999; Reich et al., 2003). Assessing nonlinear change fluctuations and covariation warrants investigation with intensive longitudinal data (Ram & Grimm, 2009). Such research can enhance understanding of the multifaceted, dynamic pathways linking affective well-being and longevity.

Conclusions

This study demonstrates that maintained and enhanced trajectories of affective well-being, encompassing PA and NA, are tied to greater survival in older adults. The results underscore the utility of modeling multidimensional affective patterns over time in relation to longevity (Ong & Steptoe, 2020; Zaninotto et al., 2016). Jointly assessing shifts in PA and NA provided insights into their interconnected influences on mortality risk. Overall, the findings reveal heterogeneous pathways linking dynamic affective states to survival, aligned with theoretical perspectives on socioemotional aging (Carstensen et al., 1999; Charles, 2010). Deteriorating affective trajectories marked by declining PA and rising NA carried the highest mortality risk and may warrant targeted intervention. In contrast, improving and flourishing trajectories signaled more adaptive patterns tied to lower mortality. This research highlights the value of applying person-centered analytical techniques to elucidate multidimensional affective pathways related to health outcomes.

References


